

Some Properties of the Cantor Distribution

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Abstract

The Cantor distribution is defined as a random series

$$\frac{1-\vartheta}{\vartheta} \sum_{i \geq 1} X_i \vartheta^i,$$

where ϑ is a parameter and the X_i are random variables that take the values 0 and 1 with probability $1/2$. The moments and order statistics are discussed, as well as a “Fibonacci” variation. Connections to certain trees and splitting processes are also mentioned.

1. Cantor distribution

1.1. **Random series.** The Cantor distribution with parameter ϑ ($0 < \vartheta \leq 1/2$) was introduced in [5] by the random series

$$X = \frac{\bar{\vartheta}}{\vartheta} \sum_{i \geq 1} X_i \vartheta^i,$$

where the X_i are independent with the distribution $\Pr[X_i = 0] = \Pr[X_i = 1] = \frac{1}{2}$, and $\bar{\vartheta} = 1 - \vartheta$. The name stems from the special case $\vartheta = \frac{1}{3}$, since then this process gives exactly those numbers from the interval $[0, 1]$ that have a ternary expansion solely consisting of the digits 0 and 2. We might alternatively consider an infinite (random) word $w_1 w_2 \dots$ over the alphabet $\{0, 1\}$ and a map **value**, defined by

$$\text{value}(w_1 w_2 \dots) = \frac{\bar{\vartheta}}{\vartheta} \sum_{i \geq 1} w_i \vartheta^i.$$

1.2. **Moments of the distribution.** We abbreviate $a_n = \mathbb{E}[X^n]$. The aim is to solve the recursion formula (from [5])

$$a_n = \frac{1}{2(1-\vartheta^n)} \sum_{k=0}^{n-1} \binom{n}{k} \bar{\vartheta}^{n-k} \vartheta^k a_k, \quad a_0 = 1.$$

Let us introduce the exponential generating function $A(z) = \sum_{n \geq 0} a_n \frac{z^n}{n!}$. The functional equation involving $A(z)$, once solved by iteration, gives

$$A(z) = \prod_{k \geq 0} \frac{1 + e^{\bar{\vartheta} \vartheta^k z}}{2}.$$

In order to derive an asymptotic equivalent of a_n , the Poisson generating function $B(z) = e^{-z}A(z)$ has to be considered. Using ‘‘Mellin’’ techniques to derive an asymptotic expansion of $\log B(z)$ when z tends to infinity and a ‘‘de-poissonization’’ argument which suggests the approximation $a_n \sim B(n)$, one gets

$$E[X^n] = a_n = F(\log_{1/\vartheta} n) n^{-\log_{1/\vartheta} 2} \left(1 + O\left(\frac{1}{n}\right) \right).$$

The function $F(x)$ is periodic of period 1 and has known Fourier coefficients. The mean of $F(x)$ is for instance

$$-\frac{1}{2 \log \vartheta} \int_0^\infty \prod_{k \geq 1} \frac{1 + e^{-\bar{\vartheta} \vartheta^k z}}{2} e^{-\bar{\vartheta} x} x^{\log_{1/\vartheta} 2 - 1} dx.$$

1.3. Order statistics. Let us consider n random independent variables Y_1, \dots, Y_n from a Cantor distribution. The average value $E[\min(Y_1, \dots, Y_n)]$ of the smallest value among them is denoted by a_n . The coefficients a_n obey the following recursion

$$(2^n - 2\vartheta)a_n = \bar{\vartheta} + \vartheta \sum_{k=1}^{n-1} \binom{n}{k} \vartheta a_k.$$

Considering now not exactly the Poisson generating function $A(z) = \sum_{k \geq 0} a_n \frac{z^n}{n!}$ but rather

$$\hat{A}(z) = \frac{1}{e^z - 1} A(z) = \sum_{n \geq 0} \hat{a}_n \frac{z^n}{n!},$$

a simpler equation can be obtained. Indeed, one has

$$\hat{A}(2z) = \vartheta \hat{A}(z) + \frac{\bar{\vartheta}}{e^z + 1}.$$

The coefficients \hat{a}_n can be extracted directly from this equation (equating coefficients of $\frac{z^n}{n!}$ on both sides). Going back to the original coefficients a_n , we have the explicit solution

$$a_n = -\bar{\vartheta} \sum_{k=0}^{n-1} \binom{n}{k} \frac{B_{k+1}}{k+1} \frac{2^{k+1} - 1}{2^k - \vartheta},$$

where B_n denotes a Bernoulli number. An approach based on Rice’s method finally gives an asymptotic equivalent of a_n

$$a_n \sim n^{\log_2 \vartheta} \frac{2\vartheta - 1}{\vartheta \log 2} (\Gamma(-\log_2 \vartheta) \zeta(-\log_2 \vartheta) + \delta(\log_2 n)),$$

where $\zeta(s)$, $\Gamma(s)$ and $\delta(s)$ denote respectively the Riemann’s zeta function, the gamma function and a periodic function with period 1 and a very small amplitude (provided ϑ is not too close to 0).

2. Cantor-Fibonacci distribution

2.1. Fibonacci restriction. The Cantor distribution might be viewed as a mapping value over a set of random words over a binary alphabet. We might also think about *restricted words*, according to the *Fibonacci restriction*, that two adjacent letters ‘1’ are not allowed. The set of (finite) Fibonacci words \mathcal{F} is given by

$$\mathcal{F} = \{0, 01\}^* \{\epsilon + 1\}.$$

In the original setting (*Cantor distribution*) probabilities are simply introduced by saying that each letter of $\{0, 1\}$ can appear with probability $\frac{1}{2}$. Here the situation is more complicated. We say

that each word of Fibonacci of length m is equally likely. There are F_{m+2} such words, with F_{m+2} denoting the $(m+2)$ th Fibonacci number. As an example, consider the classical Cantor case with $\vartheta = \frac{1}{3}$ and $m = 3$. Then the values

$$\text{value}(000) = 0, \quad \text{value}(001) = \frac{2}{27}, \quad \text{value}(010) = \frac{2}{9}, \quad \text{value}(100) = \frac{2}{3}, \quad \text{value}(101) = \frac{20}{27}$$

appear, each with probability $\frac{1}{5}$. The generating function $F(z)$ of Fibonacci words, according to their lengths is easily derived from the definition of \mathcal{F} above,

$$F(z) = \frac{1+z}{1-z-z^2} = \sum_{m \geq 0} F_{m+2} z^m.$$

Note that

$$F_n = \frac{1}{\sqrt{5}} (\alpha^n - \beta^n) \quad \text{with} \quad \alpha = \frac{1+\sqrt{5}}{2} \quad \text{and} \quad \beta = \frac{1-\sqrt{5}}{2}.$$

2.2. Moments of the Cantor-Fibonacci distribution. Let us consider the generating functions

$$G_n(z) := \sum_{w \in \mathcal{F}} (\text{value}(w))^n z^{|w|},$$

where $|w|$ denotes the length of the Fibonacci word w . The quantity

$$\frac{[z^m]G_n(z)}{[z^m]F(z)}$$

is the n th moment, when considering words of length m . Then we let m tend to infinity to get a limit called M_n (note that taking limits wasn't necessary for the independent original case). The recursion for **value**, when restricted to Fibonacci words, is

$$\begin{aligned} \text{value}(0w) &= \vartheta \cdot \text{value}(w) \\ \text{value}(10w) &= \bar{\vartheta} + \vartheta^2 \cdot \text{value}(w). \end{aligned}$$

These formulae translate almost directly to generating functions according to the recursive definition $\mathcal{F} = \epsilon + 1 + \{0, 10\}\mathcal{F}$. Thus it gives an explicit recursion formula for the functions $G_n(z)$

$$G_n(z) = \frac{1}{1 - \bar{\vartheta}^n z - \vartheta^{2n} z^2} \left[\bar{\vartheta}^n z + z^2 \sum_{i=0}^{n-1} \binom{n}{i} \bar{\vartheta}^{n-i} \vartheta^{2i} G_i(z) \right].$$

Since we only consider the limit for $m \rightarrow \infty$, we can get the asymptotic behaviour noting that both $F(z)$ and $G_n(z)$ have the same dominant singularity at $z = 1/\alpha$ and also that it is a simple pole. Consequently, we have (due to a ‘‘pole cancellation’’)

$$M_n = \lim_{m \rightarrow \infty} \frac{[z^m]G_n(z)}{[z^m]F(z)} = \lim_{z \rightarrow 1/\alpha} \frac{G_n(z)}{F(z)}.$$

Therefore we have the following theorem

Theorem 1. *The moments of the Cantor-Fibonacci distribution fulfill the following recursion: $M_0 = 0$ and for $n \geq 1$*

$$M_n = \frac{1}{\alpha^2 - \alpha \vartheta^n - \vartheta^{2n}} \sum_{i=1}^n \binom{n}{i} \bar{\vartheta}^{n-i} \vartheta^{2i} M_i.$$

2.3. The asymptotic behaviour of the moments. A rough estimate shows that $M_n \approx \lambda^n$. We might infer that $\lambda = \bar{\vartheta} + \lambda\vartheta^2$, so that $\lambda = \frac{1}{1+\bar{\vartheta}}$. It is not rigorous but we can set

$$m_n := M_n \cdot (1 + \vartheta)^n$$

anyway and show that this sequence has nicer properties. As before the recurrence on the coefficients m_n and then the exponential generating function $m(z) = \sum_n m_n \frac{z^n}{n!}$ need to be considered. Finally the Poisson transformed function $\hat{m}(z) = e^{-z}m(z)$ obeys the functional equation

$$\hat{m}(z) = \frac{e^{-\bar{\vartheta}z}}{\alpha} \hat{m}(\vartheta z) + \frac{1}{\alpha^2} \hat{m}(\vartheta^2 z).$$

Because $m_n \sim \hat{m}(n)$, the next step considers the behaviour of $\hat{m}(z)$ for $z \rightarrow \infty$. Using the Mellin transform (and the Mellin inversion formula), we have the following theorem

Theorem 2. *The n th moment M_n of the Cantor-Fibonacci distribution has for $n \rightarrow \infty$ the following asymptotic behaviour*

$$M_n = (1 + \bar{\vartheta})^{-n} \Phi(-\log_{\vartheta} n) n^{\log_{\vartheta} \alpha} \left(1 + O\left(\frac{1}{n}\right) \right),$$

where $\Phi(x)$ is a periodic function with period 1 and known Fourier coefficients. The mean (zereth Fourier coefficient) is given by

$$-\frac{1}{\log \vartheta} \int_0^{\infty} \frac{e^{-\bar{\vartheta}z}}{\alpha} \hat{m}(\vartheta z) z^{-\log_{\vartheta} \alpha - 1} dz.$$

Note that here, $\frac{e^{-\bar{\vartheta}z}}{\alpha} \hat{m}(\vartheta z)$ is merely considered as an auxiliary function. This integral can be computed numerically by replacing $\hat{m}(\vartheta z)$ by the first few values of its Taylor expansion, which can be obtained through the recursion formula on the coefficient m_n . As an example, the classical case $\vartheta = \frac{1}{3}$ gives (apart from small fluctuations),

$$M_n \sim .6160498 n^{-.4380178} 0.75^n.$$

The fact that in an asymptotic formula the generating function itself, evaluated at a certain point, appears is not at all uncommon in combinatorial analysis.

References

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